

# A compensation election for binary social choice

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**We report experimental results examining the properties of a bidding mechanism, the “Compensation Election,” which is designed to implement a simple binary choice between two options. We may think of the group decision problem as a choice between a new rule and the status quo. The rule and the status quo are each common outcomes that apply across all individuals, but the value or cost that they induce on each individual differs according to each individual’s circumstances: some gain, some lose, and others are unaffected by a change to the new from the old. Rather than casting votes, each subject submits a bid reflecting his willingness to pay to induce the group to select one option and the amount he wishes to be paid if the alternative option is selected. The Compensation Election chooses the option that receives the highest sum of bids. We find that, although the Compensation Election allows subjects to strategically bid above their value (or even for the option they do not prefer), such behavior is not the norm. We also find that subjects’ bids more truthfully reveal their values when there are more bidders in the election.**

experimental economics | public goods | auctions | voting

**W**icksell is credited with the insight that, in principle, if a proposal yields an outcome that is common to all members of a collective, and that outcome provides a preferred social state for the members of the collective, then there exists a compensation scheme whereby those who gain from the outcome can compensate those who lose and the outcome can command unanimous consent (1). This is Wicksell’s “principle of unanimity and voluntary consent in taxation.” For example, if there is a public proposition to alter a local property zoning restriction and the total capital gains accruing to its supporters exceed the total capital losses incurred by its detractors, then the value produced by altering the restriction is sufficient to compensate those who suffer a loss. With the compensation, all are made better off and the proposition is unanimously approved.

## Prior Research

Incentive-compatible methods for solving public-good decision problems of this kind were provided as early as 1971 (2–4). Vickery had earlier specified a mechanism for solving the market monopoly problem but which also could be applied to a public good (5). These solutions were computational schemes for charging each decision maker the opportunity cost that his decision (bid) imposed on all others.

All these mechanisms were thought by Vickery and others to be too complex for practical application. But several laboratory and field experimental studies of simplified versions of these mechanisms suggest that practical solutions might be devised (6–8).

A small, more recent literature has developed around Varian’s Compensation Mechanism (9), a procedure for inducing decision makers in small groups to make socially efficient decisions by requiring them to compensate those harmed by the decisions. The mechanism relies on the standard theoretical assumption that all participants know the values and costs of all other participants. By contrast, the mechanism considered in this article is designed for environments in which there is uncertainty about the preferences of other participants. Varian’s Compens-

ation Mechanism has been experimentally studied several times (10, 11), but it is the incomplete information condition that is of practical importance and that has been most studied in experimental market economies.

The Unanimity Principle was implemented in laboratory experiments with a procedure (the “Auction Election”) similar to the one we use here in that value information was private and subjects tendered bids indicating how much they would pay (or need to be paid) to implement a given choice (12). In six experiments, five groups arrived at the socially efficient outcome.

In the experimental design we report here, two major changes are made in implementing the above Auction Election. First, each decision is dichotomous: A or B. We may think of it as the standard single issue voting proposition where A is the proposal and B is the status quo. Second, each election is a one-shot decision with no in-round feedback of information (although subjects participate in a large number of elections sequentially.) We will refer to this procedure as the Compensation Election.

## Experimental Design and Research Questions

**Environment.** In the Compensation Election subjects make decisions that determine which of two mutually exclusive public goods will be chosen. Every subject  $i \in I$  is assigned a private value for each good  $k \in \{A, B\}$  in each round  $j \in \{1, \dots, 36\}$ . For simplicity, subjects’ values were normalized on their preferred good so that the value for their preferred good,  $c$ , is  $v_{ij}^c \geq 0$ , and their value for their less preferred good is  $v_{ij}^{-c} = 0$  with  $c \in \{A, B\}$ . Thus, a subject always has a value of 0 for one good and a non-negative value for the other. Our design models a case in which  $v_{ij}^k \sim N(\mu_k, \sigma_k^2)$ .

In our experimental environment, a subject’s value for his preferred good is determined by the net value  $v_{ij}^N \equiv v_{ij}^A - v_{ij}^B$ , which is distributed  $N(\mu_N \equiv \mu_A - \mu_B, \sigma_N^2 \equiv \sigma_A^2 + \sigma_B^2)$ . A subject’s value for his preferred good is given by  $v_{ij}^c = |v_{ij}^N|$ . Informally, a subject’s preferred good is A if  $v_{ij}^A > v_{ij}^B$  and B if  $v_{ij}^A < v_{ij}^B$ .

Our design consists of two treatment variables: the size of the cohort,  $I$ , and the distribution out of which the  $v_{ij}^N$  are drawn. Cohorts in the large and small treatments consisted of 18 subjects and 6 subjects, respectively. Each cohort participated in 36 elections. An election could be one of three election types: Even, Close, or Landslide. Each of the 3 types occurred 6 times in the first 18 elections and 6 times in the next 18 elections, and the ordering of the elections was randomized within these blocks.

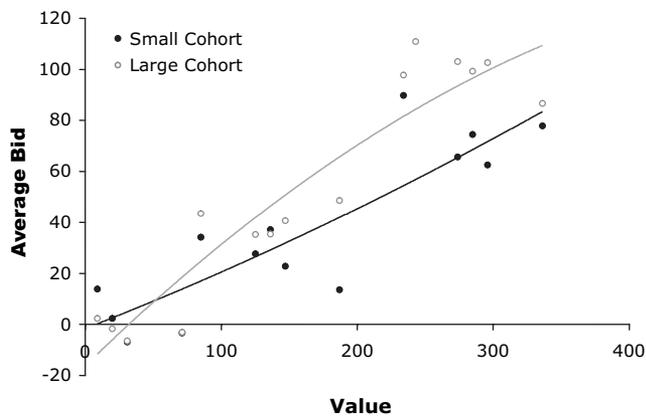
Election types are differentiated by the distributions of  $v_{ij}^N$ , which are common knowledge. Specifically, although the variance of  $v_{ij}^N$  is held constant throughout all elections, the mean is varied from election type to election type. Because we are studying small samples but wish to study behavior in cases in which the asymptotic properties of the distribution hold (as in a large-scale public decision or election), the draws we use in our

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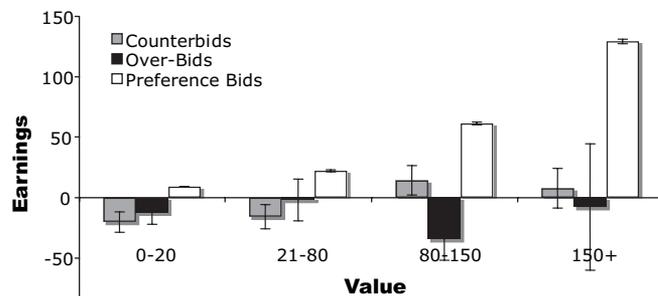
**Fig. 1.** Average bid by value and cohort size. The lines are quadratic fits to the data under each cohort size. ●, Small cohort; ○, large cohort.

**Results**

In the results reported below, we make use of data only from the last 18 rounds (the last half) of each experimental session. Our results, therefore, reflect the behavior of subjects experienced in bidding under the Compensation Election.

Our research questions concern the characteristics of the empirical bid function: the observed relationship between subjects' bids and their values. We use two techniques to study this mapping. First, Fig. 1 shows average bids at each value induced in the experiment for large cohorts and for small cohorts. Fig. 1 also includes a quadratic fit for each of these treatments to allow for curvature in our visual assessment of the bid function.

Second, to statistically pose the questions raised in *Research Questions*, we estimate the bid function using a random effects feasible generalized least squares regression model reported in Table 1. The model estimates the linear relationship between subjects' bids and their values. The regressors include (i) the intercepts made up of the interactions between treatment indicator variables for cohort size and election type, shown in the first six rows of Table 1, and (ii) slope terms estimated for cohort size and



**Fig. 2.** Average earnings by bid type over several ranges of values with standard error bars. Stippled bars, counter bids; black bars, over-bids; white bars, preference bids.

election type dummy variables interacted with the subjects' values, listed in the next six rows in Table 1. The result is separate estimates of intercepts and the marginal effect of values on bids for each treatment. We also include each of these treatment indicator variables interacted with an additional indicator variable, taking a value of 1 for observations in which the bidder prefers the losing or socially disfavored outcome. These interactions allow us to look for the type of off-preference bidding effect discussed in *Research Question 4*. Our model includes a subject-specific random (intercept) term. Although we allow for heteroskedasticity across subjects, we impose homoskedasticity within subjects.

Because counterbids are equivalent to negative bids on preferred goods, we code them as such in the data used to generate both Fig. 1 and our statistical results.

**Result 1.** There is a direct positive relationship between bids and values in all treatments and the bid function is roughly linear. The first bolded section of Table 1 contains slope term estimates for the bid function for all treatments. In all treatments, bids increase in value, and the coefficient is statistically significantly greater than zero. Fig. 1 shows the mean bids at each value under each cohort size with quadratic curves fit to the data. These quadratic fits show little curvature under either cohort size.

**Table 1. Results from mixed effects model estimating the bid function**

Treatment condition	Coefficient	SE	z	P value
Large × Even	-1.2818	2.9346	-0.44	0.662
Small × Even	-0.4190	4.5396	-0.09	0.926
Large × Close	-4.9018	4.3538	-1.13	0.26
Small × Close	-8.1903	6.7350	-1.22	0.224
Large × Landslide	-5.6806	3.5306	-1.61	0.108
Small × Landslide	-10.0253	5.4615	-1.84	0.066
<b>Large × Close × Value</b>	<b>0.4521</b>	<b>0.0227</b>	<b>19.9</b>	<b>0.000</b>
<b>Small × Close × Value</b>	<b>0.3784</b>	<b>0.0351</b>	<b>10.77</b>	<b>0.000</b>
<b>Large × Even × Value</b>	<b>0.4646</b>	<b>0.0160</b>	<b>28.92</b>	<b>0.000</b>
<b>Small × Even × Value</b>	<b>0.3451</b>	<b>0.0248</b>	<b>13.89</b>	<b>0.000</b>
<b>Large × Landslide × Value</b>	<b>0.4296</b>	<b>0.0180</b>	<b>23.87</b>	<b>0.000</b>
<b>Small × Landslide × Value</b>	<b>0.3152</b>	<b>0.0278</b>	<b>11.32</b>	<b>0.000</b>
Loser × Large × Close	3.6997	5.8799	0.63	0.529
Loser × Small × Close	8.2610	9.0957	0.91	0.364
Loser × Large × Landslide	9.8330	7.9780	1.23	0.218
Loser × Small × Landslide	11.4155	12.3093	0.93	0.354
<b>Loser × Large × Close × Value</b>	<b>0.0098</b>	<b>0.0321</b>	<b>0.31</b>	<b>0.759</b>
<b>Loser × Small × Close × Value</b>	<b>-0.0150</b>	<b>0.0496</b>	<b>-0.3</b>	<b>0.762</b>
<b>Loser × Large × Landslide × Value</b>	<b>0.0375</b>	<b>0.0448</b>	<b>0.84</b>	<b>0.403</b>
<b>Loser × Small × Landslide × Value</b>	<b>0.0726</b>	<b>0.0685</b>	<b>1.06</b>	<b>0.289</b>

Large and Small, indicator variables indicating the cohort size; Close, Even, and Landslide, indicator variables for the election type.

**Result 2.** The empirical bid function does not admit important off-preference bid behavior and subjects exhibit a strong tendency to submit preference bids. The first unbolded section of Table 1 shows intercept terms for each treatment. These intercept estimates are all insignificantly different from zero. Because, by *Result 1*, bids are increasing in value, the bid function does not, on average, feature negative bids at any value. Because counterbids are negative bids, the empirical bid function shows insignificant counterbidding.

The first bolded section of Table 1 contains estimates of the slope terms in the bid function. These slope terms are all significantly smaller than one, meaning subjects do not, on average, over-bid. Because there is neither counterbidding nor over-bidding in the empirical bid function, we may conclude that off-preference bidding is atypical behavior in our data.

Furthermore, we note that this bidding behavior is individually rationally sustainable, given the aggregate outcomes that obtain in our data. Off-preference bidding yields lower expected returns than preference bidding.

Fig. 2 shows the average earnings resulting from counterbids, over-bids, and preference bids over several value ranges. At low values, off-preference bidding results in negative expected returns, whereas preference bidding guarantees positive returns. Moreover, preference bids provide significantly higher returns at all levels of value.

**Result 3.** Bidding is affected by cohort size, but unaffected by the election type. In particular, subjects tend to bid a significantly larger proportion of their value in large cohorts than in small cohorts.

We conduct pair-wise Wald tests comparing intercepts and slope terms across treatments. Intercepts are statistically indistinguishable across treatments. Slope terms are also statistically indistinguishable across election types. However, slope terms under large cohorts are significantly larger than slope terms in small cohorts in Even ( $P = 0.001$ ) and Landslide elections ( $P = 0.0006$ ) and marginally larger under Close elections ( $P = 0.0784$ ). Table 1 shows that the estimated slope term in a large cohort is 0.12 points higher than in a small cohort in Even elections, 0.11 points higher in Landslide elections, and 0.08 higher in Close elections.

**Result 4.** Bids by those not preferring the socially preferred good are indistinguishable from bids by those preferring the efficient good. The second unbolded and bolded portions of Table 1 contain intercept and slope terms for the bid function interacted with an indicator variable that takes a value of 1 only if the bidder prefers the minority good.<sup>\*\*</sup> These terms measure the marginal effect on bidding behavior of those not preferring the socially preferred good. The coefficients on these terms are insignificantly different from zero in every case.

## Discussion

We experimentally tested a bidding mechanism for referendum decision making that allows participants to protect their assets from confiscation by a majority rule winning coalition. Overall, the empirical bid function in the Compensation Election is qualitatively

similar to the bid function under many commonly studied private auctions. Empirically, there is a direct relationship between values and bids, no tendency for bidders to bid over their value, a tendency for bidders to bid on the good they prefer, and a tendency for bidders to bid less than their value (i.e., to shade). This latter property is theoretically less than optimal in public goods provision, but we note that majority rule suffers from precisely the same defect and can potentially be more inefficient. Moreover, despite its pervasive popularity, majority rule voting often violates democratic principles of freedom and fair play in confiscating minority asset value to the benefit of a majority.

Our experiments were designed to evaluate subject bidding behavior under the compensation features of our mechanism. Therefore, our parameter choices were not suited for systematic efficiency comparisons of the Compensation Election with the majority rule mechanism. Future research could return to this issue, and to the important question of how robust our findings are for much larger collectives than 18 people. Our findings that 18-subject cohorts shade bids to a lesser degree than their 6-subject counterparts are suggestive, but far from conclusive.

In closing we note that there are many variations on the mechanism suggested by applied problems in historical practice. For example, the mechanism can be studied for application to classic “holdout” problems in acquiring right-of-way for railroads, power lines or pipelines, and land assemblages for private or public projects (airports, arenas, shopping malls, etc.) in which any sequential acquisition and information-release procedure can invite one or a few holdout owners who seek to capture a disproportionate share of the surplus. Eminent domain as a solution to this problem has itself led to spectacular and unpopular abuses. An advantage of the bidding mechanism we propose is that all decisions are made simultaneously; all parties are confronted by the same uncertainty regarding the bids of their counterparts; and, where more than one site or right-of-way path is feasible, the efficiency of the common outcome decision can benefit from contestability among the alternatives.

## Materials and Methods

Experiments were conducted in the Interdisciplinary Center for Economic Science laboratory at George Mason University. We recruited 108 undergraduate subjects through an online recruiting system. Ten sessions were conducted. At 4 of these sessions, 18 subjects participated, and, in 6 of them, 6 subjects participated. Recruited subjects had no experience in the experiment and were not allowed to participate in more than one of the experiments reported here. However, all recruited subjects had previous experience with the double auction institution and were thus familiar with laboratory procedures. Sessions were held between September 2003 and October 2004, and each lasted  $\approx 2$  h.

In a typical session, subjects waited in a reception room until a sufficient number had arrived. Twelve to 18 subjects were taken into a laboratory room and seated at computers behind partitions so that they could not see each other’s screens. A strict no-talking rule was enforced as soon as subjects entered the laboratory. At each session, subjects were given identical computerized instructions that gave detailed information on the environment and rules of the institution and gave subjects experience with the interface. Subjects earned, on average, \$33.22 in small-cohort sessions and \$32.78 in large-cohort sessions.

<sup>\*\*</sup>Because there are no minority goods in the Even election, there are no such interaction terms for Even election treatments.

1. Wicksell K (1896/1967) in *Classics in the Theory of Public Finance*, eds Musgrave RA, Peacock AT (St Martin’s, New York), pp 72–118.
2. Clarke E (1971) *Public Choice* 11:17–33.
3. Groves T (1973) *Econometrica* 41:617–633.
4. Groves T, Ledyard J (1977) *Econometrica* 45:783–809.
5. Vickery W (1961) *J Finance* 16:8–37.
6. Bohm P (1972) *Euro Econ Rev* 3:111–130.
7. Scherr B, Babb E (1975) *Public Choice* 23:35–48.
8. Ferejohn J, Noll R (1976) *Am Econ Rev* 66:267–273.
9. Varian H (1994) *Am Econ Rev* 84:1278–1293.
10. Andreoni J, Varian H (1999) *Proc Natl Acad Sci USA* 96:10933–10938.
11. Chen Y, Gazzale R (2004) *Am Econ Rev* 94:1505–1535.
12. Smith V (1977) *J Polit Econ* 85:1125–1140.
13. Downs A (1957) *An Economic Theory of Democracy* (HarperCollins, New York).
14. Milgrom PR, Webber RJ (1982) *Econometrica* 50:1089–1122.